## 1 Language Models

With a language model, we can estimate how likely a string is English (or how likely that an English speaker would have uttered the string). Naturally, we want an SMT system not only to contain precise translations but also combine them into fluent English

Language models in general are functions, that take a string and return an estimate of how likely is that string is a proper English phrase.

LMs help us with reordering, for example:

$$
p_{\mathrm{LM}}(\text { the house is small })>p_{\mathrm{LM}}(\text { small the is house })
$$

Also, they help with the choice of the right word:

$$
p_{\mathrm{LM}}(\mathrm{I} \text { am going home })>p_{\mathrm{LM}}(\mathrm{I} \text { am going house })
$$

As a rule, the functions are estimated on large monolingual (much more abundant than parallel) corpora of target text.

$$
\begin{aligned}
p\left(w_{1}^{k}\right) & =p\left(w_{1}, w_{2}, \ldots, w_{k}\right) \\
& =p\left(w_{1}\right) p\left(w_{2} \mid w_{1}\right) p\left(w_{3} \mid w_{1}, w_{2}\right) \ldots p\left(w_{k} \mid w_{1}, \ldots w_{k-1}\right)
\end{aligned}
$$

It is natural to further assume, that only the previous history of $n-1$ words (for some $n$ ) matters for predicting the next word. We are going to use maximum likelihood estimation, so another reason for limiting history is that sufficiently long phrases don't appear in any fixed-size corpus, therefore we must break the computation of the language model into smaller steps.

$$
\begin{aligned}
p\left(w_{1}^{k}\right) & =p\left(w_{1}, w_{2}, \ldots, w_{k}\right) \\
& =p\left(w_{1}\right) p\left(w_{2} \mid w_{1}\right) p\left(w_{3} \mid w_{1}, w_{2}\right) \ldots p\left(w_{i} \mid w_{i-n+1}, \ldots w_{i-1}\right) \ldots p\left(w_{k} \mid w_{k-n+1}, \ldots w_{k-1}\right)
\end{aligned}
$$

This is an application of the chain rule for conditional probabilities:

$$
\begin{aligned}
p(x \mid y) & =\frac{p(x, y)}{p(y)} \\
p(y) p(x \mid y) & =p(x, y)
\end{aligned}
$$

## 1.1 n-gram Language Models

### 1.1.1 Definition of n-gram

The Markov assumption tells us, that only $n-1$ words in history matter.

$$
\begin{equation*}
p\left(w_{1}, \ldots w_{k}\right) \simeq \prod_{i=1}^{k} p\left(w_{i} \mid w_{i-n+1}^{i-1}\right) \tag{1}
\end{equation*}
$$

bigram: $n=2 \Rightarrow$ First order Markov model
Example: $p\left(w_{1}, \ldots, w_{k}\right) \simeq p\left(w_{1}\right) p\left(w_{2} \mid w_{1}\right) \ldots p\left(w_{k} \mid w_{k-1}\right)$
trigram: $n=3 \Rightarrow$ Second order Markov model
Example: $p\left(w_{1}, \ldots, w_{k}\right) \simeq p\left(w_{1}\right) p\left(w_{2} \mid w_{1}\right) p\left(w_{3} \mid w_{1}, w_{2}\right) \ldots p\left(w_{k} \mid w_{k-2}, w_{k-1}\right)$

### 1.2 Estimating n-gram probabilities

$$
p\left(w_{i} \mid w_{i-n+1}^{i-1}\right)=\frac{\operatorname{count}\left(w_{i-n+1}^{i}\right)}{\operatorname{count}\left(w_{i-n+1}^{i-1}\right)}
$$

Maximum likelihood estimation:

$$
\begin{align*}
p\left(w_{2} \mid w_{1}\right) & =\frac{\operatorname{count}\left(w_{1}, w_{2}\right)}{\operatorname{count}\left(w_{1}\right)}  \tag{2}\\
p\left(w_{k} \mid w_{k-1}, w_{k-2}\right) & =\frac{\operatorname{count}\left(w_{k-2}, w_{k-1}, w_{k}\right)}{\operatorname{count}\left(w_{k-2}, w_{k-1}\right)}
\end{align*}
$$

Example from a 3-gram LM trained on the European Parliament proceedings:

| the green (total: |  |  |
| :--- | ---: | ---: |
| Word | Count | Prob. |
| paper | 801 | 0.458 |
| group | 640 | 0.367 |
| light | 110 | 0.063 |
| party | 27 | 0.015 |
| ecu | 21 | 0.012 |


| the red (total: 225) |  |  |
| :--- | ---: | ---: |
| Word | Count | Prob. |
| cross | 123 | 0.547 |
| tape | 31 | 0.138 |
| army | 9 | 0.040 |
| card | 7 | 0.031 |
| , | 5 | 0.022 |


| the blue (total: 54 ) |  |  |
| :--- | ---: | ---: |
| Word | Count | Prob. |
| box | 16 | 0.296 |
| flag | 6 | 0.111 |
| , | 6 | 0.111 |
| angel | 3 | 0.056 |

### 1.3 Fit quality: Perplexity

Problem: How well does a language model perform/fit the data? (e.g., what order is enough? which smoothing technique is better?)

One possible measure is Perplexity, which is defined based on cross-entropy: Entropy \& perplexity:

$$
\begin{aligned}
H(p(x)) & =-\mathbb{E}_{p(x)} \log _{2} p(x) \\
& =-\sum_{x} p(x) \log _{2} p(x) \\
P P(p(x)) & =2^{H(p(x))}
\end{aligned}
$$

Now, entropy of a language is defined as a limit of per-word entropy:

$$
H(L, p(x))=-\lim _{n \rightarrow \infty} \frac{1}{n} \sum_{\forall w_{1}^{n}} p\left(w_{1}^{n}\right) \log p\left(w_{1}^{n}\right),
$$

where $w_{1}^{n}$ is a $n$-word sequence from the language $L$.

## Shannon-McMillan-Breimann Theorem:

$$
H(L, p(x)) \simeq-\frac{1}{n} \log p\left(w_{1}^{n}\right), \text { for large } n
$$

Thus, a sufficiently long single sequence is representative for the whole language. We can now compute the perplexity of model using a (large) corpus.

In our case the independent events are the observed n-grams, $w_{i-n+1}^{i}$, from sentence $s$. Therefore, the expectation is taken over the emprical distribution:

$$
\begin{aligned}
H\left(p_{L M}(s)\right) & =-\frac{1}{|s|-n+1} \log _{2} p_{L M}(s) \\
& =-\frac{1}{|s|-n+1} \sum_{i}^{|s|-n+1} \log _{2} p\left(w_{i} \mid w_{i-n+1}^{i-1}\right) . \\
P P & =2^{H\left(p_{L M}(s)\right)}
\end{aligned}
$$

The idea is that a model with smaller Perplexity on unseen data is better (model is "less surprised" to see this new data).

Example (for the sentence "I would like to commend the rapporteur on this work"):

| Prediction | $p_{\text {LM }}$ | $-\log _{2} p_{\text {LM }}$ |
| :---: | :---: | :---: |
| $p_{\text {LM }}(i\|</ \mathrm{s}><\mathrm{s}\rangle$ ) | 0.109 | 3.197 |
| $p_{\text {LM }}($ would $\mid<$ s $>$ I) | 0.144 | 2.791 |
| $p_{\text {LM }}$ (like\|i would) | 0.489 | 1.031 |
| $p_{\text {LM }}$ (to\|would like) | 0.905 | 0.144 |
| $p_{\text {LM }}$ (commend\|like to) | 0.002 | 8.794 |
| $p_{\text {LM }}$ (the\|to commend) | 0.472 | 1.084 |
| $p_{\text {LM }}$ (rapporteur\|commend the) | 0.147 | 2.763 |
| $p_{\text {LM }}$ (on\|the rapporteur) | 0.056 | 4.150 |
| $p_{\text {LM }}$ (his\|rapporteur on) | 0.194 | 2.367 |
| $p_{\text {LM }}$ (work\|on his) | 0.089 | 3.498 |
| $p_{\text {LM }}$ (. \|his work) | 0.290 | 1.785 |
| $p_{\text {LM }}(</ \mathrm{s}>\mid$ work .) | 0.99999 | 0.000014 |
|  | Average | 2.634 |

Example (comparison of different $n$-gram models):

| Word | Unigram | Bigram | Trigram | 4-gram |
| :--- | ---: | ---: | :---: | :---: |
| i | 6.684 | 3.197 | 3.197 | 3.197 |
| would | 8.342 | 2.884 | 2.791 | 2.791 |
| like | 9.129 | 2.026 | 1.031 | 1.290 |
| to | 5.081 | 0.402 | 0.144 | 0.113 |
| commend | 15.487 | 12.335 | 8.794 | 8.633 |
| the | 3.885 | 1.402 | 1.084 | 0.880 |
| rapporteur | 10.840 | 7.319 | 2.763 | 2.350 |
| on | 6.765 | 4.140 | 4.150 | 1.862 |
| his | 10.678 | 7.316 | 2.367 | 1.978 |
| work | 9.993 | 4.816 | 3.498 | 2.394 |
| l | 4.896 | 3.020 | 1.785 | 1.510 |
| </s> | 4.828 | 0.005 | 0.000 | 0.000 |
| Average | 8.051 | 4.072 | 2.634 | 2.251 |
| Perplexity | 265.136 | 16.817 | 6.206 | 4.758 |

### 1.4 Smoothing

High-order $n$-grams are not very frequent in corpora. Zero probabilities of $n$-grams destroy our computation of sentence probabilities, since it will be floored to zero.
Smoothing attempts to fill in missing statistics in order to avoid zero probabilities (any product of the form (1) is zero if at least one factor is zero).

### 1.4.1 Laplace smoothing

## "Add-one"-smoothing

$$
\begin{array}{ll}
\text { unsmoothed: } & p=\frac{c}{n}, \\
\text { "Add-one"-smoothing: } p=\frac{c+1}{n+v}, & v=\text { size of vocabulary } \tag{4}
\end{array}
$$

## "Add $\alpha$ "-smoothing

$$
\begin{equation*}
p=\frac{c+\alpha}{n+\alpha v}, \quad \alpha<1, \alpha \text { optimized on held-out set } \tag{5}
\end{equation*}
$$

Example: 2-grams in Europarl

| Count | Adjusted count |  | Test count |
| :---: | :---: | :---: | :---: |
| $c$ | "add 1" | "add $\alpha$ | $t_{c}$ |
| 0 | 0.00378 | 0.00016 | 0.00016 |
| 1 | 0.00755 | 0.95725 | 0.46235 |
| 2 | 0.01133 | 1.91433 | 1.39946 |
| 3 | 0.01511 | 2.87141 | 2.34307 |
| 4 | 0.01888 | 3.82850 | 3.35202 |
| 5 | 0.02266 | 4.78558 | 4.35234 |
| 6 | 0.02644 | 5.74266 | 5.33762 |
| 8 | 0.03399 | 7.65683 | 7.15074 |
| 10 | 0.04155 | 9.57100 | 9.11927 |
| 20 | 0.07931 | 19.14183 | 18.95948 |

$\alpha=0.00017$
$t_{c}=$ average count of $n$-gram in test set that occured $c$ times in training corpus

### 1.4.2 "Deleted estimate"-smoothing

$N_{r}$ is the count of $n$-grams with training count $r$,
$T_{r}$ is the count in test data of $n$-grams with training count $r$.
We also have an estimate $r^{*}=\frac{T_{r}}{N_{r}}$.
Then we switch train and test and combine the results:

$$
\begin{equation*}
r^{*}=\frac{T_{r}^{1}+T_{r}^{2}}{N_{r}^{1}+N_{r}^{2}} \tag{6}
\end{equation*}
$$

This is similar to 2 -fold cross validation.

| Count <br> $r$ | Count of counts $N_{r}$ | Count in held-out $T_{r}$ | Exp. count $E[r]=T_{r} / N_{r}$ | Test count $t_{C}$ |
| :---: | :---: | :---: | :---: | :---: |
| 0 | 7,515,623,434 | 938,504 | 0.00012 | 0.00016 |
| 1 | 753,777 | 353,383 | 0.46900 | 0.46235 |
| 2 | 170,913 | 239,736 | 1.40322 | 1.39946 |
| 3 | 78,614 | 189,686 | 2.41381 | 2.34307 |
| 4 | 46,769 | 157,485 | 3.36860 | 3.35202 |
| 5 | 31,413 | 134,653 | 4.28820 | 4.35234 |
| 6 | 22,520 | 122,079 | 5.42301 | 5.33762 |
| 8 | 13,586 | 99,668 | 7.33892 | 7.15074 |
| 10 | 9,106 | 85,666 | 9.41129 | 9.11927 |
| 20 | 2,797 | 53,262 | 19.04992 | 18.95948 |

### 1.4.3 "Good-Turing"-smoothing

Derivation Question: what is the probability of observing n-gram "in the wild" (i.e. in some test test), given that we saw the n-gram with some frequency in the training set?

Assume we have a sample $S$ of size $|S|$ drawn from the true probability ngram distribution $p(w)$. Define $c(w)$ to be the number of times n-gram $w$ occurs in $S$. For integer $r \geq 0$ let $S_{r}=\{w: c(w)=r\}$. For example, $S_{0}$ is the set of n-grams not seen in $S$.

Define a random variable $M_{r}$ (depending on $S$ ) to be the probability of drawing an n-gram in the set $S_{r}$, that is $M_{r}=\sum_{w \in S_{r}} p(w)$. For example, $M_{0}$ is the "missing mass" - total probability of words not occuring in the sample $S$.

Had we known $M_{r}$, the true probability $P\left(w \mid w \in S_{r}\right)$ of drawing again some word $w \in S_{r}$ would be

$$
\begin{equation*}
\frac{M_{r}}{\left|S_{r}\right|}, \tag{7}
\end{equation*}
$$

i.e., total mass divided by total number of distinct elements.

Example: To see the need for a smoothing, imagine we have sampled a large $S$ where each n-gram occurs exactly once (quite unlikely event). The naive way of estimating $M_{1}$ would be

$$
\frac{\# \text { of times } w \text { occurs in } S \times \# \text { of different words we are ok with }}{\text { total size of } S}=\frac{k \times\left|S_{1}\right|}{|S|}=\frac{1 \times|S|}{|S|}=1
$$

However, for any reasonable distribution $p(w)$ the probability $M_{1}$, given such an unlikely sample $S$, should be close to 0 .

Let us find the expectation of $M_{r}$ :

$$
\begin{aligned}
\mathbb{E}\left[M_{r}\right] & =\sum_{w} p(w) P\left[w \in S_{r}\right] \\
& =\sum_{w} p(w)\binom{|S|}{r} p^{r}(w)(1-p(w))^{|S|-r} \\
& =\sum_{w}\binom{|S|}{r} p^{r+1}(w)(1-p(w))^{|S|-(r+1)}(1-p(w)) \frac{r+1}{r+1} \\
& =\sum_{w}\binom{|S|}{r+1} p^{r+1}(w)(1-p(w))^{|S|-(r+1)}(1-p(w)) \frac{r+1}{|S|-r} \\
& =\sum_{w} P\left[w \in S_{r+1}\right](1-p(w)) \frac{r+1}{|S|-r} \\
& =\frac{r+1}{|S|-r} \sum_{w} P\left[w \in S_{r+1}\right]-\frac{r+1}{|S|-r} \sum_{w} p(w) P\left[w \in S_{r+1}\right] \\
& =\frac{r+1}{|S|-r} \sum_{w} p(w)\left|S_{r+1}\right|-\frac{r+1}{|S|-r} \mathbb{E}\left[M_{r+1}\right] \\
& =\frac{r+1}{|S|-r} \mathbb{E}\left[\left|S_{r+1}\right|\right]-\frac{r+1}{|S|-r} \mathbb{E}\left[M_{r+1}\right]
\end{aligned}
$$

As $0 \leq M_{r+1} \leq 1$, if $k \ll|S|$ the last term is close to zero. Therefore for $k \ll|S|$ an almost unbiased estimate of $M_{r}$ is

$$
\frac{r+1}{|S|-r} \mathbb{E}\left[\left|S_{r+1}\right|\right] \simeq \frac{r+1}{|S|}\left|S_{r+1}\right|,
$$

where we assume that means of (large) $\left|S_{r+1}\right|$ can be estimated reliably just by $\left|S_{r+1}\right|$.

Final formula Using formula (7) we adjust the actual counts $r$ to expected counts $r^{*}$ with this formula:

$$
\begin{equation*}
r^{*}=(r+1) \frac{N_{r+1}}{N_{r}} \tag{8}
\end{equation*}
$$

where $N_{r}=\left|S_{r}\right|$ is the number of n-grams that occur exactly $r$ times in our corpus and $N_{0}$ is the total number of $n$-grams not occuring in the corpus.
In practice, although very well justified, the GT estimate is not alone for n-gram smooothing, because the formula may give very noisy estimates for large $r$, where the $n_{r+1}$ statistics not reliable or absent $(=0)$, resulting in suboptimal performance.

| Count <br> $r$ | Count of counts <br> $N_{r}$ | Adjusted count <br> $r^{*}$ | Test count <br> $t$ |
| :--- | ---: | ---: | ---: |
| 0 | $7,514,941,065$ | 0.00015 | 0.00016 |
| 1 | $1,132,844$ | 0.46539 | 0.46235 |
| 2 | 263,611 | 1.40679 | 1.39946 |
| 3 | 123,615 | 2.38767 | 2.34307 |
| 4 | 73,788 | 3.33753 | 3.35202 |
| 5 | 49,254 | 4.36967 | 4.35234 |
| 6 | 35,869 | 5.32928 | 5.33762 |
| 8 | 21,693 | 7.43798 | 7.15074 |
| 10 | 14,880 | 9.31304 | 9.11927 |
| 20 | 4,546 | 19.54487 | 18.95948 |

### 1.5 Back-Off and Interpolation

## Back-Off:

In a given corpus, we may never observe a collocation like "Scottish beer drinkers" or "Scottish beer eaters", because they both have a count of 0 . Therefore our smoothing methods will assign them the same probability.

A better idea: back-off to bigrams, like "beer drinkers" and "beer eaters".

## Interpolation:

Higher and lower order n-gram models have different strengths and weaknesses:

- high-order n-grams are sensitive to more context, but have sparse counts
- low-order n-grams consider only very limited context, but have robust counts

Combine them:

$$
\begin{aligned}
p_{I}\left(w_{3} \mid w_{1}, w_{2}\right)= & \lambda_{1} p_{1}\left(w_{3}\right) \\
& +\lambda_{2} p_{2}\left(w_{3} \mid w_{2}\right) \\
& +\lambda_{3} p_{3}\left(w_{3} \mid w_{1}, w_{2}\right)
\end{aligned}
$$

With a lot of training data, we can trust the higher order language models more and assign them higher weights. We require that:

$$
\begin{aligned}
\forall \lambda_{n}: & 0 \leq \lambda_{n} \leq 1 \\
\sum_{n} \lambda_{n} & =1
\end{aligned}
$$

## Recursive Interpolation (Jelinek-Mercer smoothing)

Recursive definition of interpolation:

$$
\begin{aligned}
p_{n}^{I}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right) & =\lambda_{w_{i-n+1}, \ldots, w_{i-1}} p_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right) \\
& +\left(1-\lambda_{w_{i-n+1}, \ldots, w_{i-1}}\right) p_{n-1}^{I}\left(w_{i} \mid w_{i-n+2}, \ldots, w_{i-1}\right)
\end{aligned}
$$

Training $\lambda_{w_{i-n+1}, \ldots, w_{i-1}}$ can be done with the EM algorithm, however, partitioning them into buckets according to $c\left(w_{i-n+1}^{i-1}\right)$ and using the same $\lambda$ for all counts in the same buckets has a good balance of quality and eficiency.

## Recursive Back-Off (Katz smoothing)

To fix the GT smoothing we would trust the highest order language model that contains an n-gram:
$p_{n}^{B O}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right)=\left\{\begin{array}{c}d_{n}\left(w_{i-n+1}, \ldots, w_{i-1}\right) p_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right) \\ \text { if } \operatorname{count}_{n}\left(w_{i-n+1}, \ldots, w_{i}\right)>k \\ \alpha_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right) p_{n-1}^{B O}\left(w_{i} \mid w_{i-n+2}, \ldots, w_{i-1}\right) \\ \text { otherwise }\end{array}\right.$
The constant $k$ is usually set to be a small integer, often zero or found emprically by cross-validation.

First note that for GT smoothing the total probability of lost probability mass due to smoothing and that is distributed among unseen n-grams is equal to $n_{1} / N$, where $N=\sum_{r>0} r n_{r}$, with the contribution of each $n$-gram with count $r$ to be equal to $\left(1-d_{r}\right) \frac{r}{N}$.

Consider two cases:

1. $k=0$

- as the GT estimation reduces counts (and estimated probabilites) we can just use the discount values GT smoothing gives as a discounting function $d_{n}\left(w_{1}, \ldots, w_{n-1}\right)$.
- coefficients $\alpha_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right)$ then just collect all the missing probability mass:

$$
\begin{aligned}
& \beta\left(w_{i-n+1}, \ldots, w_{i-1}\right) \equiv 1-\sum_{\operatorname{count}_{n}\left(w_{i-n+1}, \ldots, w_{i}\right)>k} d_{n}\left(w_{i-n+1}, \ldots, w_{i-1}\right) p_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right) \\
& \alpha_{n}\left(w_{i} \mid w_{i-n+1}, \ldots, w_{i-1}\right)=\frac{\beta\left(w_{i-n+1}, \ldots, w_{i-1}\right)}{\sum_{\operatorname{count}_{n}\left(w_{i-n+1}, \ldots, w_{i}\right) \leq k} p_{n-1}^{B O}\left(w_{i} \mid w_{i-n+2}, \ldots, w_{i-1}\right)}
\end{aligned}
$$

2. $k>0$

- set $d_{r}$ for all $n$-grams occuring more than $k$ times to be equal to 1 (i.e., considered to be a reliable estimate)
- otherwise require that contributions $d_{r}$ of seen n-grams are propotional to the GT contributions

$$
1-d_{r}=\mu\left(1-r^{*} / r\right)
$$

- still the total distributed mass as dictated by GT smoothing should be left unchanged

$$
\sum_{r>0} n_{r}\left(1-d_{r}\right) \frac{r}{N}=\frac{n_{1}}{N}
$$

- can be shown that the solution is

$$
d_{r}=\frac{\frac{r^{*}}{r}-\frac{(k+1) n_{k+1}}{n_{1}}}{1-\frac{(k+1) n_{k+1}}{n_{1}}}
$$

## Recursion grounding

- 1-order model: ML (or some smoothed) unigram model
- 0 -order model: uniform model $p\left(w_{i}\right)=1 /|V|$

